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Regular Research Article

The Cognitive Profile of Older Adults With Treatment-Resistant Depression: An Analysis of the OPTIMUM Randomized Controlled Trial

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ABSTRACT

Objective: Major depressive disorder in older adults (late-life depression; LLD) is frequently associated with cognitive impairment, and some deficits (e.g., executive function) bave been associated with a higher level of treatment resistance. However, the cognitive profile of treatment-resistant LLD (TR-LLD) bas not been characterized. We hypothesized that patients with TR-LLD would show deficits in cognitive function, especially executive function, and that executive function deficits would predict poorer response to pharmacotherapy. **Design:** Secondary analysis of baseline cognitive data from OPTIMUM, a multicenter RCT evaluating pharmacotherapy strategies for TR-LLD. **Setting:** Five outpatient academic medical centers (4 US, 1 Canada). **Participants:** About 369 participants aged 60 and older from the OPTIMUM study. **Measurements:** Baseline scores on individual tasks and composite scores from the NIH Toolbox-Cognition Battery were transformed into demographically-adjusted T-scores and compared to

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published norms. Impairments in the set shifting and inhibitory control tasks were investigated as predictors of depressive symptom change following treatment using ANCOVA models. **Results:** Participants had low performance on tasks evaluating inhibitory control, processing speed, verbal/nonverbal memory, and the fluid composite, but normative performance on working memory and set shifting. Participants had high estimated premorbid IQ (superior Performance on oral reading recognition). Age and physical comorbidity negatively associated with processing speed. Impairments in set shifting predicted less improvement in depressive symptoms; impairments in inhibitory control did not. **Conclusions:** Participants with TR-LLD presented with broad cognitive deficits relative to healthy norms. Given poorer outcomes following standard pharmacotherapy associated with impaired set shifting, future research needs to identify alternative treatment strategies. (Am J Geriatr Psychiatry 2025; 33:361–371)

Highlights

- What is the primary question addressed by this study? What is the impact of treatment-resistant late-life depression (TR-LLD) on cognitive function?
- What is the main finding of this study? Individuals with TR-LLD exhibit broad-based cognitive deficits, including in some areas of executive function but not others.
- What is the meaning of the finding? Cognitive deficits in TR-LLD present a common clinical challenge that may impact on clinical course.

INTRODUCTION

M ajor depressive disorder in older adults (latelife depression; LLD) is a risk factor for cognitive decline and dementia.¹ Many patients with LLD present with some cognitive impairment, and up to 40% meet diagnostic criteria for Mild Cognitive Impairment (MCI), conferring additional risk for future dementia.² Accordingly, the cognitive function of individuals suffering from LLD is a subject of ongoing interest.

Previous studies have shown that multiple cognitive domains can be affected in LLD, including attention, information processing speed, learning and memory, and executive function.3 With respect to executive function, several studies have found that patients with LLD frequently present with specific impairments in inhibitory control ^{3,4} and set shifting.³ Comorbid executive dysfunction has been associated with poorer response to treatment with antidepressant medication,^{4,5} an effect possibly mediated by poor treatment adherence.⁶ Furthermore, executive dysfunction frequently persists even with treatment of LLD.^{7–9}

To date, there are limited published data on cognitive function in individuals with treatment-resistant LLD (TR-LLD). The cognitive profile of these patients is important to understand in part because treatment resistance is highly prevalent and may be associated with underlying neurodegenerative changes related to Alzheimer's Disease and/or vascular disease.¹⁰ Thus, cognitive impairment in TR-LLD may be a both a marker of treatment resistance and target for treatment.

In this context, we characterized domains of cognitive function in participants in the Optimizing Outcomes of Treatment-Resistant Depression in Older Adults (OPTIMUM) study.¹¹ In this analysis, we aimed to: 1) describe the cognitive profile of participants with TR-LLD across multiple cognitive domains; and 2) investigate whether baseline cognitive function predicts change in depressive symptoms

following an acute course of pharmacotherapy. We tested the hypotheses that: 1a) participants with TR-LLD would exhibit impaired cognitive function at baseline, based on published norms; 1b) participants with TR-LLD would demonstrate greatest impairment in measures of executive function, including inhibitory control and set shifting; and 2) based on our previous findings linking executive dysfunction and treatment resistance in LLD¹⁰ participants with TR-LLD and impaired performance in measures of executive function (i.e., inhibitory control and set shifting) at baseline would show less improvement in depressive symptoms following protocolized pharmacotherapy than those with TR-LLD and intact executive function. We also explored whether sociodemographic and clinical characteristics explain the associations between cognitive impairment and TR-LLD.

METHODS

OPTIMUM Study

The design and implementation of the OPTIMUM study have been described in detail elsewhere.^{11,12} Briefly, this was a pragmatic, randomized controlled trial (RCT) conducted at 5 sites (4 in the US, 1 in Canada) that compared the effectiveness of augmentation and switch strategies for TR-LLD. Participants were excluded if they showed evidence of possible dementia (based on a Short Blessed Test score >10), had unstable medical illness, or had a diagnosis of Parkinson's disease. Step 1 compared aripiprazole augmentation, bupropion augmentation, or switch to bupropion; Step 2 compared lithium augmentation to nortriptyline switch. Participants who were ineligible for Step 1 due to a previous medication trial were enrolled in Step 2 directly, while participants who did not remit after Step 1 were offered subsequent enrolment in Step 2 as a continuation of their participation. Each step lasted 10 weeks and participants completed the Montgomery-Åsberg Depression Rating Scale (MADRS; Montgomery & Åsberg¹³) at baseline and after 10 weeks. At baseline, they also completed the Patient-Reported Outcomes Measurement Information System (PROMIS) anxiety questionnaire¹⁴ and the Cumulative Illness Rating Scale-Geriatric (CIRS-G).¹⁵ Participants in the OPTIMUM trial were individuals age 60 or older with a diagnosis of major depressive disorder and treatment-resistant depression, as defined by a lack of remission of major depression after two or more antidepressant trials of adequate dose and duration.

Cognitive Measures

All participants were asked to complete the NIH Toolbox Cognition Battery, Version 2 (NIHTB-CB) at baseline, prior to beginning their initial treatment in either Step 1 or Step 2, depending on eligibility. Only Step 1 participants and those who enrolled directly in Step 2 were included in this analysis. Because, as has been previously reported,¹⁶ this trial utilized remote procedures to improve accessibility for individuals for whom transportation to the study site was not feasible, many participants were unable to complete the NIHTB-CB due to the need for on-site attendance for this instrument. This instrument was selected because of its extensive validation including available normative data, including in older adults.¹⁷ Participants were asked to complete six tasks: (1) the Dimensional Change Card Sort Test ("Card Sort task"; a set shifting measure of cognitive flexibility); (2) the Flanker Inhibitory Control and Attention Test ("Flanker task"; inhibitory control and selective attention); (3) the List Sorting Working Memory Test ("List Sorting task"; working memory); (4) the Pattern Comparison Processing Speed Test ("Pattern Comparison task"; processing speed); (5) the Picture Sequence Memory Test ("Picture Sequence task"; verbal and nonverbal episodic memory); and (6) the Oral Reading Recognition Test ("Oral Reading task"; an oral word recognition task as a surrogate of premorbid intellectual function. All cognitive testing was completed on a tablet. Version 2.1 was used for all tests. For the Picture Sequence task, three forms with different sets of stimuli were used (Forms A, B, and C) and rotated intraindividually between assessments to mitigate practice effects.

We collected baseline scores for the fluid composite performance and the 6 individual tasks of the NIH Toolbox Cognition Battery for all participants who completed these assessments. We used the fluid composite score computed by the NIH Toolbox software, which averages the normalized scores of each measure completed in a given testing session,¹⁸ with the exception of the Oral Reading Recognition task. For individual tasks, where individual component scores were available (e.g., the reaction time and accuracy scores for the Flanker task), the combination scale score was used. To characterize the sample at baseline, we used the T-scores derived from the test norms¹⁹ for each task adjusted for age, self-reported gender, and education, with a mean of 50 and a standard deviation of 10.

Statistical Analysis

Of 740 participants in the OPTIMUM study (619 from Step 1 and 121 from Step 2), we included in the analysis 369 participants (315 from Step 1 and 54 from Step 2) who completed cognitive testing at baseline. We compared mean T scores for each baseline test within our sample against population norms using single-sample Student's t tests against a standardized population mean of 50, using Bonferroni correction for multiple comparisons. Correlations of cognitive scores with sociodemographic and clinical variables of interest were screened using a univariate correlation matrix; variables that were correlated (p <0.10) with each cognitive score were included in a final multivariate linear regression model (Pearson method), together with three variables of a priori interest: baseline MADRS, PROMIS, and CIRS-G scores.

Next, we compared the change in MADRS scores following treatment in participants with versus without significant impairment (defined as a T-score of <40, i.e., at least one standard deviation below the standardized mean) in performance on two tests of executive function: the Card Sort task (set shifting) and the Flanker Task (inhibitory control). For each test, we analyzed the main effect of significant impairment in baseline performance on the change in MADRS score from baseline to the end of the treatment phase, using ANCOVA models adjusted for age, self-reported gender, education, baseline PROMIS score, baseline MADRS score, and other sociodemographic or clinical variables if they were correlated with the MADRS score change at a p-value of <0.1.

RESULTS

Summary of Included Sample

Of the 369 participants included in this analysis, the majority were female (239; 64.8%), with a mean (SD) age of 69.1 (6.9) years and a mean (SD) 14.7 (2.9) years of formal education. Table 1 presents the base-line sociodemographic and clinical characteristics of the sample. These characteristics were broadly similar

Sociodemographic	Step 1 (n = 315)	Step 2 (n = 54)	
Self-identified female gender, n (%)	203 (64.4)	36 (66.7)	
Age in years, mean (SD)	69.0 (6.4)	67.5 (4.2)	
Education in years, mean (SD)	14.8 (3.1)	15.5 (2.4)	
Married, n (%)	118 (37.4)	18 (33.3)	
Race, n (%)			
Asian	3 (1.0)	0 (0.0)	
Black	24 (7.7)	0 (0.0)	
Hawaiian/Pacific Islander	0 (0.0)	0 (0.0)	
Indigenous	0 (0.0)	0 (0.0)	
Multirace	6 (1.6)	1 (1.9)	
Other	14 (3.9)	0 (0.0)	
White	268 (85.6)	52 (96.3)	
Clinical			
MDD age of onset in years, mean (SD)	32.6 (20.4)	24.7 (15.3)	
Number of adequate medication trials, mean (SD)	2.4 (0.9)	2.5 (1.0)	
CIRS-G score, mean (SD) ^a	8.3 (4.8)	7.6 (4.9)	
PROMIS anxiety questionnaire T-score, mean (SD)	64.2 (6.9)	64.0 (7.3)	
MADRS score at baseline, mean (SD) ^a	22.6 (7.3)	22.8 (7.5)	
MADRS score at end of treatment, mean (SD)	16.2 (9.0)	17.6 (9.0)	

SD: standard deviation; MDD: major depressive disorder; CIRS-G: Cumulative Illness Rating Scale – Geriatric; PROMIS, patient-reported outcomes measurement information system; MADRS, Montgomery-Åsberg depression rating scale.

^a Significant difference between participants with and without baseline cognitive data.

						Confidence Interval	
Cognitive Test	Test Statistic	df	p-value	Mean Difference	Effect Size (Cohen's d)	Lower	Upper
Card sorting task	t = 0.39	365	0.699	0.21	0.02	-0.08	0.12
Flanker task	t = -25.04	364	< 0.001	-9.36^{a}	-1.31	-1.45	-1.17
Fluid composite	t = -10.75	288	< 0.001	-6.29^{a}	-0.63	-0.76	-0.51
List sorting task	t = -2.54	360	0.012	-1.2	-0.13	-0.24	-0.03
Pattern comparison task	t = -12.08	363	< 0.001	-8.67^{a}	-0.63	-0.75	-0.52
Picture sequence task	t = -5.68	350	< 0.001	-2.94^{a}	-0.30	-0.41	-0.20

TABLE 2.	Comparison of Baseline Scores on NI	H Toolbox Cognitive Batter	y Tasks Against Population Norms
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Notes: Test statistics are for single-sample t-tests using a standardized population mean T-score of 50 and hypothesis: \neq 50. Confidence intervals are for Cohen's d.

df, Degrees of freedom (differing values reflect different numbers of participants who completed each task).

Negative mean difference values correspond to a lower score on the task versus the standardized population mean.

^a Significant difference compared with standardized population mean for each test, based on an alpha threshold (Bonferroni) of 0.0071.

between the 369 participants included in this analysis and the 371 participants who were not included because of missing cognitive data, except that included participants had a slightly lower level of physical comorbidity (CIRS-G mean [SD]: 8.2 [4.8] versus 9.1 [4.6], t = 2.58, df = 733, p = 0.01) and lower baseline MADRS score (mean [SD]: 22.6 [7.4] versus 23.7 [7.1], t = 2.08, df = 714, p = 0.04). Similarly, participants who entered the study at Step 1 or directly at Step 2 had minimal differences in sociodemographic and clinical characteristics as well as baseline cognitive scores: the only statistically significant difference across 15 comparisons was in years of education, which was slightly higher in participants directly entering Step 2 (mean [SD]: 15.2 [2.6] years versus 14.6 [2.9] years, t = 2.14, df = 725, p = 0.03).

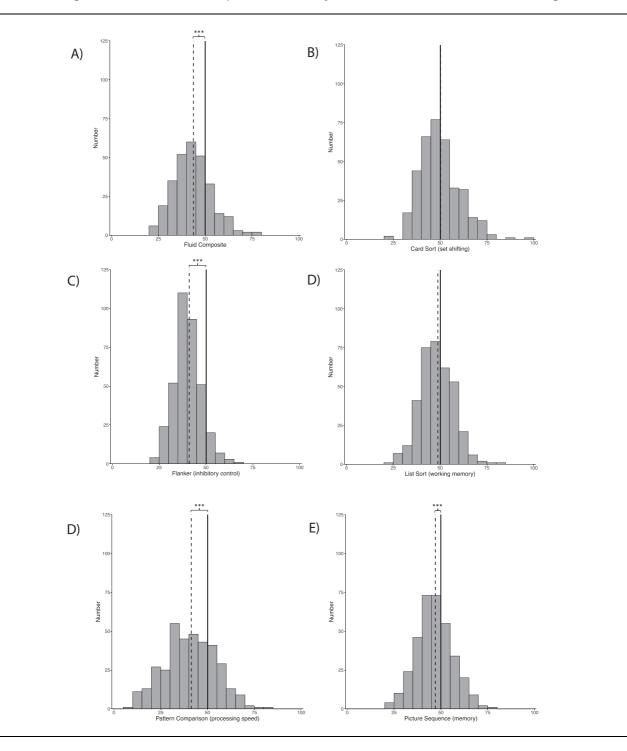
Baseline Cognitive Function

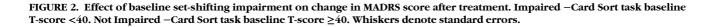
TR-LLD participants were significantly impaired on most cognitive domains at baseline, based on published demographic norms accounting for age, gender, and education (see Fig. 1A–F). A large effect size for impairment was seen in the Flanker task (inhibitory control) (mean [SD] T-score: 40.6 [7.1], t = -25.04, df = 364, p <0.001; Cohen's d = -1.31[95% CI: -1.45, -1.17]), which remained significant after correction for multiple comparisons. Moderate effect sizes for impairment were also seen on the Pattern Comparison task (processing speed) (mean [SD] T-score: 41.3 [13.7], t = -12.08, df = 363, p <0.001; Cohen's d = -0.63 [95% CI: -0.75, -0.52]) and on the Fluid Composite (mean [SD] T-score: 43.7 [9.9], t = -10.75, df = 288, p <0.001; Cohen's d = -0.63 [95% CI: -0.76, -0.51]), with a small effect size on the Picture Sequence task (memory) (mean [SD] T-score: 47.1 [9.7], t = -5.68, df = 350, p < 0.001; Cohen's d = -0.30 [95% CI: -0.41, -0.20]); all of these remained significant after correction for multiple comparisons (See Table 2). No significant impairment was seen on the List Sort task (working memory) or Card Sort task (set shifting) after correction for multiple comparisons (see Table, Supplemental Digital Content 1 for further descriptive statistics). TR-LLD participants performed significantly better than average on the oral reading comprehension task (a surrogate for premorbid intellectual function) based on published norms (mean [SD] T-score: 57.9 [8.5], t = 17.66, df = 355, p <0.001; Cohen's d = 0.94 [95% CI: 0.81, 1.06]).

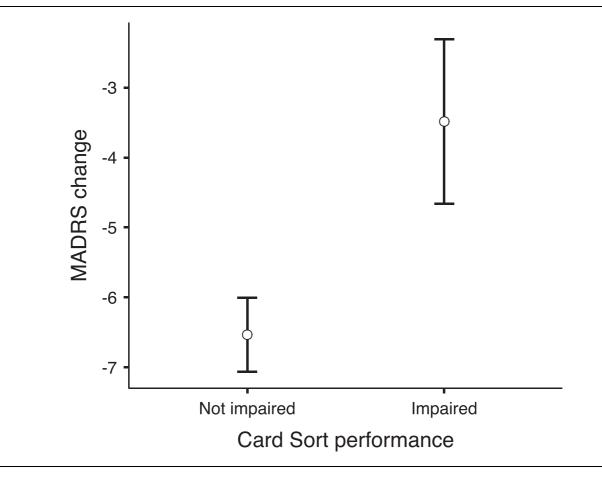
Correlates of Baseline Cognitive Function in TR-LLD

Several sociodemographic and clinical variables were correlated with baseline scores on individual cognitive tasks following univariate linear regression (see Figure, Supplemental Digital Content 2, containing the univariate correlation matrix used to screen variables). However, only two of these associations remained significant after covariate adjustment: age (estimated coefficient [unstandardized weight]: -0.377 [95% CI: -0.620, -0.134], t = -3.05, df = 313, p = 0.002) and CIRS-G scores (estimated coefficient [unstandardized weight]: -0.473 [95% CI: -0.774, -0.172], t = -3.09, df = 313, p = 0.002) each were negatively associated with scores on the Pattern Comparison task.

FIGURE 1. Distributions of cognitive scores for OPTIMUM participants at baseline. Each histogram shows the distribution of demographically-adjusted T-scores obtained in our sample from a task within the NIH Toolbox – Cognitive Battery: Fluid Composite (A), Card Sort task (B), Flanker task (C), List Sorting task (D), Pattern Comparison task (E), and Picture Sequence task (F). The standardized mean T-score for each task based on demographic norms is 50 (solid lines), with a standard deviation of 10. The mean T-score obtained in our sample for each task is indicated by a dashed line. "Significant difference from standardized mean, p <0.001.







Does Executive Dysfunction Predict Less Depressive Symptom Improvement in TR-LLD?

Of 366 participants who completed the Card Sorting task, 58 (15.8%) showed impairment, as defined by a baseline T-score on the Card Sort task of <40, while 165 of 365 participants (45.2%) showed impairment on the Flanker task. Stratification of impairment by mild (T-score: 39–30), moderate (Tscore: 29–20), or severe (T-score <20) revealed that most participants in the impaired range for each task showed mild impairment, with very few in the moderate impairment category and none in the severe category (see Table, Supplemental Digital Content 3).

Participants with impairment in set-shifting performance had less improvement in depressive symptoms following treatment than participants with better setshifting performance (adjusted mean difference in MADRS score change = -3.05 points, df = 240, p = 0.02, Cohen's d = -0.41) (see Fig. 2). Conversely, participants with impairment in inhibitory control performance (T-score of <40 on the Flanker task) did not significantly differ in change in MADRS score compared with those with higher performance. These overall effects remained significant when the analysis was restricted to Step 1 participants only.

DISCUSSION

We analyzed the cognitive profile of 369 participants with TR-LLD who were enrolled in the first step of the OPTIMUM pharmacotherapy trial. Our analysis had two main findings. First, our TR-LLD sample exhibited impairment relative to healthy older adults in composite cognition and in specific domains based on test norms, with the largest degree of impairment seen in one area of executive function, inhibitory control (as measured by the Flanker Task). However, our sample did not exhibit impairment in another measure of executive function, set shifting (as measured by the Card Sort task). These impairments were also observed in individuals who were relatively well educated and demonstrated significantly superior performance on an oral word recognition task (reflecting premorbid general intelligence). In this treatment-resistant sample, we did not observe any consistent sociodemographic or clinical correlates of baseline cognitive function, except that older age and greater medical comorbidity were associated with worse information processing speed. Second, in our analysis of baseline executive function as a predictor of depressive symptom outcomes, baseline low performance in set-shifting (but not inhibitory control) predicted less improvement in depressive symptoms following pharmacotherapy for TR-LLD.

While the cognitive profile of patients with LLD has now been well characterized,^{2,3} to our knowledge, this is the first description of the cognitive profile of patients with TR-LLD receiving pharmacotherapy. Prior work has characterized cognitive function in mid-aged adults with treatment resistant depression,^{20,21} and the cognitive profile of a sample of older adults receiving ECT has also been reported.²² Our analysis confirmed our hypothesis that participants with TR-LLD would demonstrate broad cognitive impairment at baseline, and we saw more widespread deficits than have typically been identified in non-treatment-resistant LLD.^{2,3} However, we saw a divergence in tasks measuring components of executive function, with significant impairment seen with the Flanker Task (inhibitory control) but not with the Card Sort task (set-shifting). The impairment in inhibitory control in our sample is congruent with prior findings that older depressed patients with impaired inhibitory control are less likely to respond to initial antidepressant treatment,^{4,5} which increases the likelihood that they would require further treatment as offered in the

OPTIMUM trial. The divergence we saw in baseline performance on executive function tasks underlines the complexity of executive function and the importance of considering its various putative components. Based on our results, it would be inexact to describe our participants with TR-LLD as exhibiting "executive dysfunction" broadly, and more accurate to describe them as exhibiting an impairment in inhibitory control. Taken together, these findings suggest that individuals with TR-LLD experience specific and substantial impairments in cognitive function (including in inhibitory control). They also suggest that a high premorbid intelligence ("cognitive reserve"), as estimated here by the Oral Reading task, may not be enough to mitigate this deficit; however, it should be noted that this finding requires replication with other surrogate tests for cognitive reserve, which can estimate this differently^{23,24} They also suggest that in individuals with TR-LLD, impairment is not seen exclusively in executive function; rather, cognitive impairment in this group appears to be broadly distributed among several cognitive domains, and some areas of executive function may be spared, particularly in individuals with higher premorbid intellectual functioning. This includes significant impairment in memory, the persistence of which despite adequate treatment has been previously described in the existing LLD literature, and which may suggest an increased risk of developing dementia.8

Our analysis of executive function in older patients as a predictor of response to TR-LLD pharmacotherapy yielded contrasting findings: participants with worse set-shifting experienced significantly less improvement in depressive symptoms following treatment than those without better set-shifting, while participants with worse inhibitory control experienced a similar degree of depressive symptom improvement as those with better inhibitory control. This is consistent with prior results from the IRL-GRey study²⁵ among participants who received aripiprazole augmentation following nonremission to a prospective trial of venlafaxine¹⁰ However, it contrasts with results from several studies of patients with LLD without treatment resistance: most of these studies have found an association between specific impairments in inhibitory control and less depressive symptom improvement following pharmacotherapy and this has been confirmed in a meta-analysis.⁵ Since inhibitory control predicts treatment resistance in LDD, patients with TR-LLD may have a more consistent impairment in inhibitory control such that it precludes predicting treatment response in these patients; this may lead to inhibitory control having a more homogeneous impact on treatment response in TR-LLD versus non-treatment-resistant LLD. Another factor could contribute to the different relationship between cognitive function and depressive outcomes in patients with LLD or TR-LLD: both medications used as the first treatment step of OPTIMUM, (i.e., aripiprazole and bupropion) act on dopamine; thus, they may be better suited for treating depressive symptoms in those with impaired inhibitory control than selective serotonin reuptake inhibitors or other first-line antidepressants. Given the evidence supporting a relationship between dopaminergic transmission and inhibitory control,^{26,27} future studies will need to determine whether dopaminergic medications may mitigate the deleterious impact on depressive symptom improvement attributed to executive dysfunction. Conversely, the worse outcomes experienced by the participants in our TR-LLD sample with poor set-shifting, consistent with a prior study,¹⁰ suggest that this impairment may be a trait-like feature in TR-LLD that forms part of a highly treatment-resistant subgroup. It should be noted that this subgroup was relatively small in our study, suggesting that the adverse impact of this form of executive dysfunction may be relatively uncommon even in a TR-LLD sample.

Our results should be considered in the context of some limitations. First, complete NIH Toolbox Cognition Battery data were only available for a subset of OPTIMUM participants, limiting our sample size. However, comparison of participants who did and did not have these cognitive data revealed only small differences in baseline characteristics (0.6 years of education and 1.1 points on the MADRS). It is unlikely that these small differences would have any impact on our results; additionally, all our analyses controlled for potential clinical and sociodemographic confounders. Second, we did not have a sample of patients with non-treatment-resistant LLD to which we could directly compare the cognitive profile of our participants with TR-LLD; therefore, all analyses involved either participant scores based on test norms in healthy individuals,¹⁹ or indirect reference to prior literature in non-treatment-resistant LLD samples. Third, we were limited in our ability to

compare different levels of executive function impairment on treatment outcomes, as most participants were in the mild impairment range. Fourth, our measure of set-shifting, the Dimensional Change Card Sort Test, is a complex measure that may also capture some other cognitive abilities such as inhibition and processing speed^{28–30}; however, our results are consistent with a prior study using a different measure of set-shifting (i.e., the Trail-Making Test Part B). Finally, our protocol utilized a single cognitive battery (the NIH Toolbox) and selected antidepressant agents, so these findings may not generalize to other samples utilizing different cognitive tests or medications.

CONCLUSION

Cognitive impairment is a major clinical challenge in LLD. Our results confirm that cognitive impairment is substantial and widespread in patients with TR-LLD, with multiple domains affected, including some but not all measures of executive functioning. Our results also suggest that while poor baseline set-shifting predicts less improvement in depressive symptoms following pharmacotherapy for TR-LLD, poor inhibitory control may not decrease the likelihood of responding to medications that affect the dopaminergic system. In summary, our results provide a novel description and analysis of the cognitive profile of patients with TR-LLD, and they provide insights regarding management of these patients when they present with executive dysfunction.

DISCLOSURES

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AUTHOR CONTRIBUTIONS

NJA contributed to study design, analysis, manuscript drafting and revision; HO contributed to study design, analysis, and manuscript revision; HL, DMB, PJB, JFK, EJL, MPM, BGP, and CFR contributed to study design and manuscript revision; EL contributed to data collection, analysis, and manuscript revision; JPM contributed to analysis and manuscript revision; BHM contributed to study design, manuscript drafting and revision.

DATA STATEMENT

Due to confidentiality and funding agency constraints, these data cannot be made publicly available.

PREVIOUS PRESENTATION

These data have not been previously presented orally or by poster at scientific meetings.

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*Except for the first three authors and the last author, all the authors contributed equally and are listed in alphabetical order.

SUPPLEMENTARY MATERIALS

Supplementary material associated with this article can be found in the online version at https://doi.org/10.1016/j.jagp.2024.09.018.

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